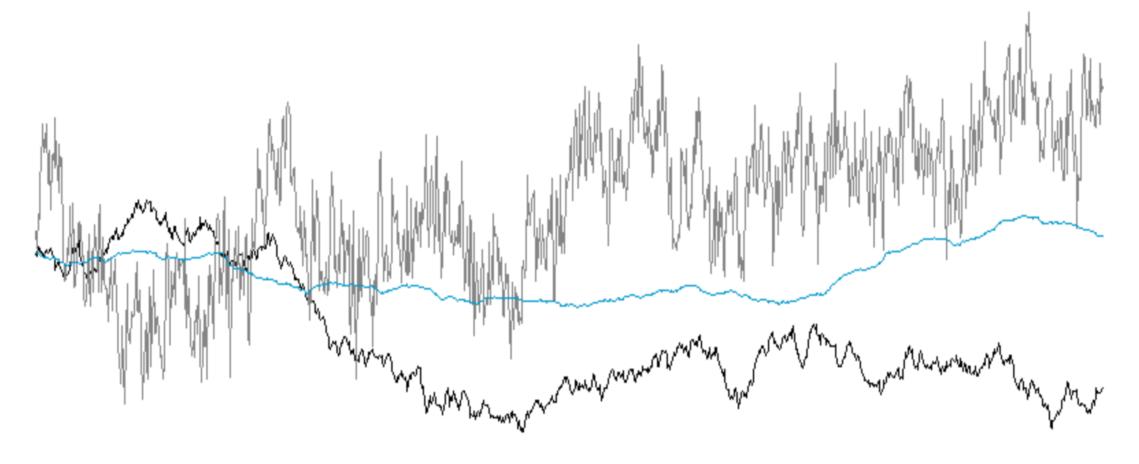
# ROUGH STOCHASTIC VOLATILITY MODELING AND ITS IMPACT ON LONG-TERM LIFE INSURANCE PRICING

IA|BE PRIZE 2021 - DUPRET JEAN-LOUP



# ROUGH STOCHASTIC VOLATILITY MODELING

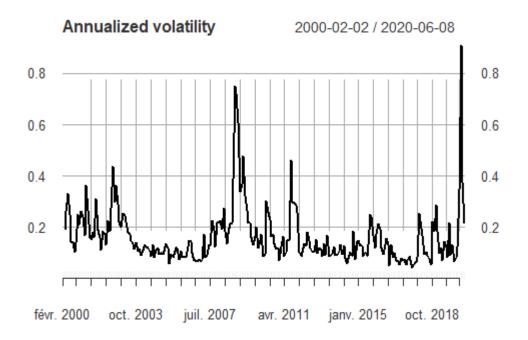
- Classical stochastic volatility models
- Fractional Brownian motion
- RFSV model
  - rBergomi
  - Rough Heston
- Calibration
- Life insurance contract

# **VOLATILITY IS NOT CONSTANT**

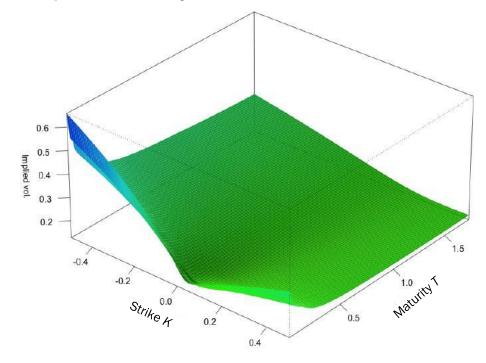
Observed volatility of stock returns is not constant but varies randomly with time:

→ Black & Scholes model inaccurate!

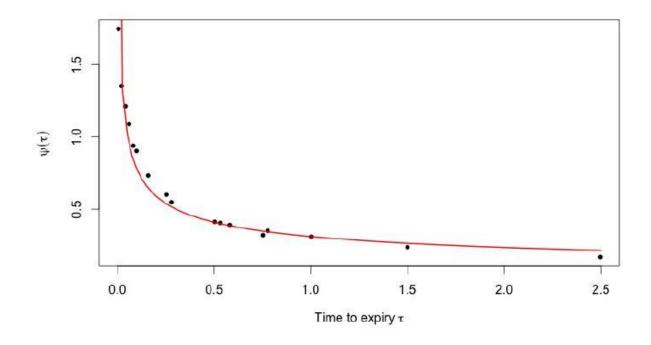
#### - Historical volatility S&P500:



#### - Implied volatility S&P500:



# **ATM VOLATILITY SKEWS: POWER-LAW**



- $\rightarrow$  In red, power-law function:  $\psi(\tau) = A \tau^{-0.4}$  .
- → In **black**, ATM volatility skew estimated for the S&P500, 6th July

### **HESTON AND BATES MODELS**

• Under a risk-neutral measure  $\mathbb{P}^*$ , the **Heston model** is given by:

$$dS_t = S_t r dt + S_t \sqrt{v_t} dW_t^*$$
  
$$dv_t = \kappa(\eta - v_t) dt + \theta \sqrt{v_t} d\hat{W}_t^*$$

with  $v_t$  the variance process and  $W^*$ ,  $\widehat{W}^*$  two correlated Brownian motions under  $\mathbb{P}^*$ .

• Under  $\mathbb{P}^*$ , the **Bates model** is given by:

$$\frac{dS_t}{S_t} = (r - \lambda \xi) dt + \sqrt{v_t} dW_t^* + (Y_t - 1) dN_t$$
$$dv_t = \kappa (\eta - v_t) dt + \theta \sqrt{v_t} d\hat{W}_t^*$$

with  $N_t \sim Poi(\lambda t)$  and  $\log(Y_t) \sim N(\mu_j, \sigma_j^2)$ 

#### **HESTON AND BATES MODELS**

#### **ADVANTAGES:**

- Incorporate mean-reverting stochastic volatility.
- Characteristic function in closed-form -> fast and efficient calibration.

#### **DRAWBACKS**:

- Implied volatility not realistic under these two models.
- Cannot reproduce the memory properties of the observed historical volatility.

# ROUGH STOCHASTIC VOLATILITY MODELING

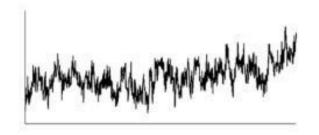
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# FRACTIONAL BROWNIAN MOTION (fBm)

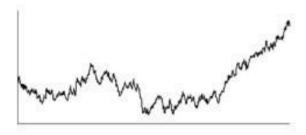
A fractional Brownian motion  $(B_t^H)_{t\geq 0}$  is a Gaussian process iif :

$$Cov(B_t^H, B_s^H) = \mathbb{E}(B_t^H B_s^H) = \frac{1}{2} \{ t^{2H} + s^{2H} - |t - s|^{2H} \} \mathbb{E} ((B_1^H)^2)$$

- Depends on the parameter  $H \in (0,1)$ , called the **Hurst index**.
- Stationarity of increments.
- Increments are **positively** correlated if H > 1/2, **negatively** correlated if H < 1/2 and **independent** if H = 1/2:



Fractional Bm : H < 1/2



Classical Bm : H = 1/2



Fractional Bm : H>1/2

## FRACTIONAL BROWNIAN MOTION - INCREMENTS

- The process of increments of the fBm  $\Delta B_t^H = B_t^H B_{t-1}^H$  is said to have :
  - Long memory for H > 1/2
  - Short memory for H < 1/2
  - → One-to-one correspondance with the regularity of fBm trajectories.

- $\Delta B_t^H$  is called a **fractional Gaussian noise** 
  - → Basis of fractional mean-reverting process (RFSV model).

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# ROUGH FRACTIONAL STOCHASTIC VOLATILITY MODEL (RFSV)

• The RFSV model is based on a fractional mean-reverting process for the log-volatility with H < 1/2:

$$dS_t = r S_t dt + \sigma_t S_t dW_t^*$$
  
$$dX_t = \lambda(\eta - X_t) dt + \nu dB_t^H$$

where  $X_t = \log \sigma_t$  and  $dB_t^H$  a fractional Gaussian noise with short memory.

• The volatility  $\sigma_t = \exp(X_t)$  is the **unique stationary solution with short memory** given by :

$$\sigma_t = \exp(X_t) = \exp\left\{\eta + \int_{-\infty}^t e^{-\lambda(t-u)} dB_u^H\right\}$$

# **RFSV - HISTORICAL VOLATILITY (CAC40)**

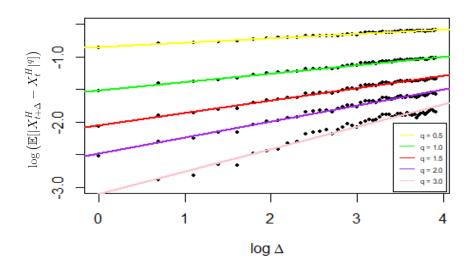
• Gatheral et al. (2014) show with  $\lambda \to 0$ :

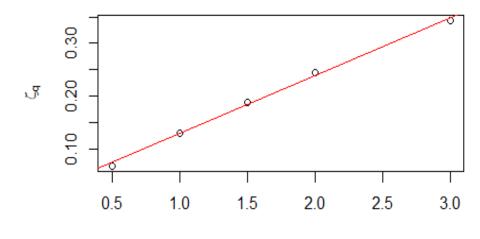
$$\mathbb{E}\Big[\sup_{t\in[0,T]}|X_t^H - X_0^H - \nu B_t^H|\Big] \to 0$$

and:

$$\mathbb{E}[|X_{t+\Delta}^H - X_t^H|^q] \to \nu^q K_q \Delta^{qH}$$

- When  $\lambda \to 0$ , the log-volatility process of the RFSV behaves as a fBm and approximately reproduces their scaling property.
- → Confirmed **empirically** with the CAC40



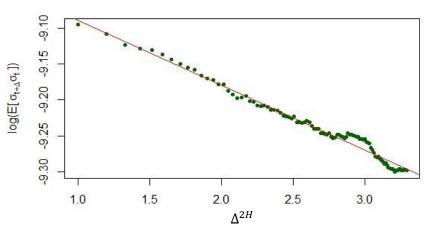


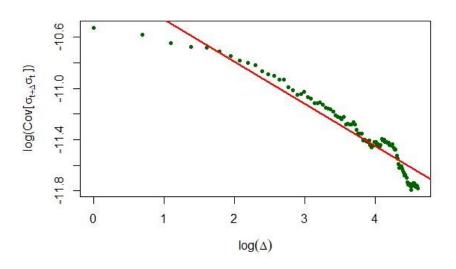
# **RFSV - HISTORICAL VOLATILITY (CAC40)**

The autocovariance of  $\sigma_t$  when  $\lambda o 0$  is given by :

$$\mathbb{E}[\sigma_{t+\Delta}\sigma_t] = \mathbb{E}[e^{X_t^H + X_{t+\Delta}^H}] \approx e^{2\mathbb{E}[X_t^H] + 2\operatorname{Var}[X_t^H]}e^{-\nu^2\frac{\Delta^{2H}}{2}}$$

- $\rightarrow \log (\mathbb{E}[\sigma_{t+\Delta}\sigma_t])$  is linear in  $\Delta^{2H}$ , which is confirmed empirically.
- $\rightarrow \mathbb{E}[\sigma_{t+\Delta}\sigma_t]$  does not behave as a **power-law function**. Nor the empirical data nor the RFSV exhibit long-term memory.



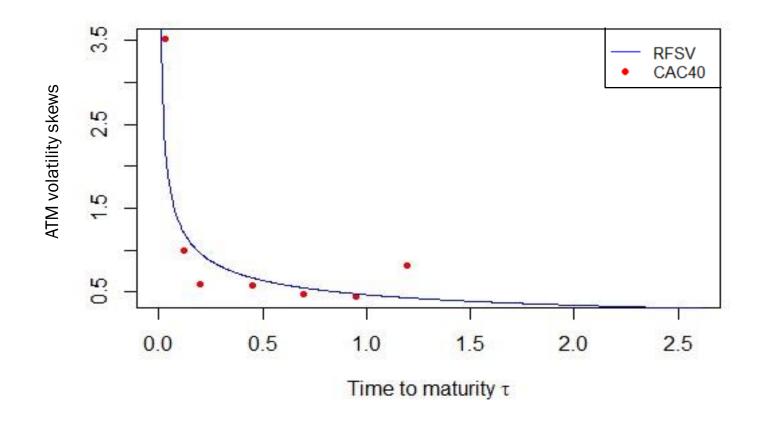


# **RFSV - HISTORICAL VOLATILITY**

The RFSV model is extremely consistent with the **observed historical volatility** due to its <u>short memory</u> and <u>rough sample paths</u> (H < 1/2).

# **RFSV - IMPLIED VOLATILITY**

→ Extremely consistent with **implied volatility** and especially with the term structure of **ATM** volatility skews :



#### **RFSV - DRAWBACK**

 <u>BUT</u>, the RFSV model is too slow for pricing and calibration since it requires a lot of slow and unstable <u>Monte-Carlo</u> simulations.

- Two more efficient models derived from the RFSV :
  - → rBergomi model
  - → Rough Heston model

#### **rBERGOMI MODEL**

• Model obtained from the RFSV by setting  $\lambda = 0$ :

$$S_T = S_t \exp\left(r(T-t) - \frac{1}{2} \int_t^T v_u \, du + \int_t^T \sqrt{v_u} \, dW_u^{*,S}\right)$$

$$v_u = \mathbb{E}^{\mathbb{P}^*}[v_u | \mathcal{F}_t] \exp\left\{\eta \sqrt{2H} \int_t^u \frac{1}{(u-s)^{1/2-H}} dW_s^* - \frac{\eta^2}{2} (u-t)^{2H}\right\} = \mathbb{E}^{\mathbb{P}^*}[v_u | \mathcal{F}_t] \mathcal{E}\left(\eta \, \tilde{W}_t^*(u)\right)$$

- A bit more efficient and stable than the RFSV model but still **not optimal** for calibration.
- The volatility generated by the rBergomi is **not stationary** since  $\lambda = 0$ .

→ Inappropriate for long-term life insurance pricing

#### **ROUGH HESTON MODEL**

Extension of the classical Heston model with a rough fractional Gaussian noise (H <  $\frac{1}{2}$ ):

$$dS_t = S_t r dt + S_t \sqrt{v_t} dW_t^*$$

$$v_t = \xi_0(t) + \frac{1}{\Gamma(H+1/2)} \int_0^t \frac{\nu}{(t-s)^{1/2-H}} \sqrt{v_s} d\hat{W}_s^*$$

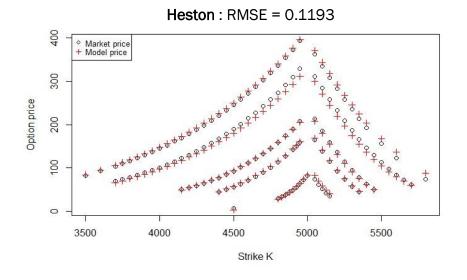
The rough Heston is excellent for pricing long-term life insurance contracts

- Stationary volatility generated by the rough Heston model.
- Only 3 parameters and a characteristic function in closed-form
  - → Pricing and calibration far more efficient and stable.
- Highly consistent with historical and implied volatility.

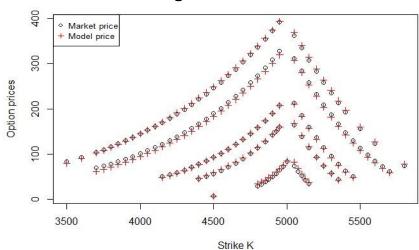
# ROUGH STOCHASTIC VOLATILITY MODELING

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# **COMPARISON OF MODEL CALIBRATIONS** (CAC 40, RMSE)



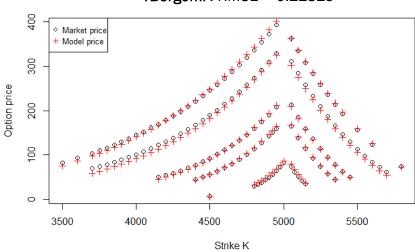
Rough Heston: RMSE = 0.0929



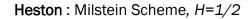
Bates: RMSE = 0.0863

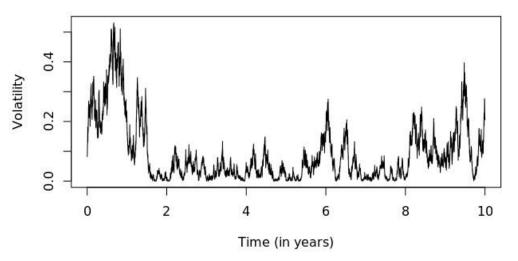
rBergomi: RMSE = 0.11315

Strike K

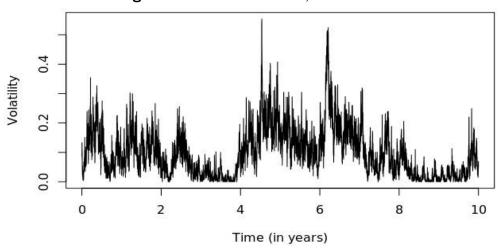


## **COMPARISON OF VOLATILITY SAMPLE PATHS**

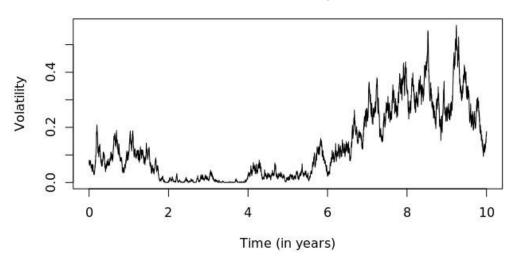




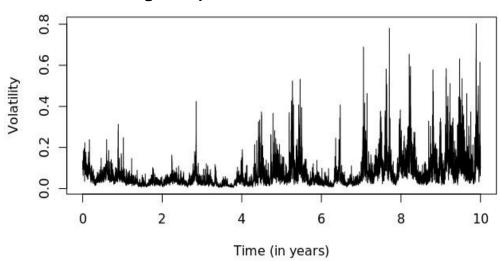
Rough Heston: Euler Scheme, H=0.123



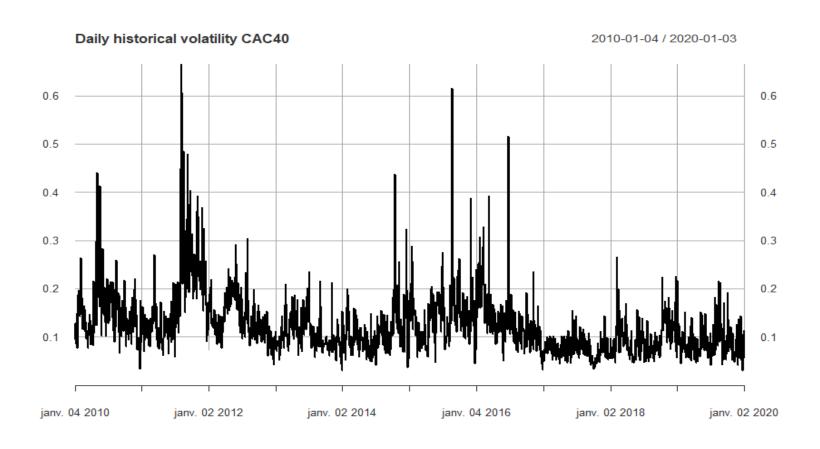
**Bates**: Milstein Scheme, H = 1/2



**rBergomi**: Hybrid Scheme, *H*=0.150



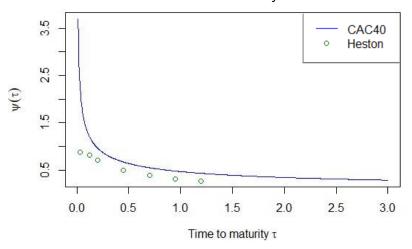
#### **OBSERVED HISTORICAL VOLATILITY CAC40**



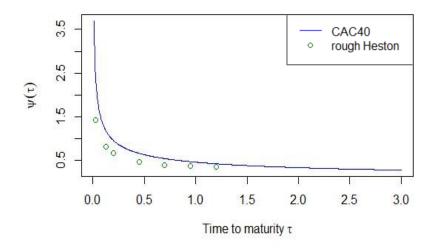
→ Visually, same volatility sample paths as rough models.

## **COMPARISON OF ATM VOLATILITY SKEWS**

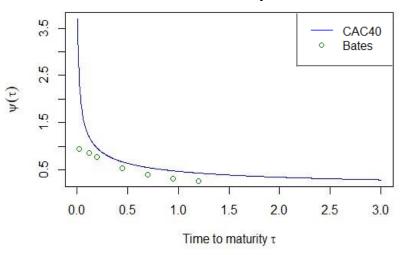
**Heston:** ATM volatility skews



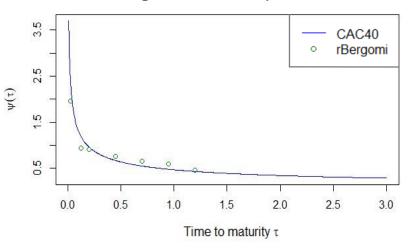
Rough Heston: ATM volatility skews



Bates: ATM volatility skews



rBergomi: ATM volatility skews



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# **EQUITY-LINKED LIFE INSURANCE CONTRACT**

- **Endowment insurance** with maturity T where the benefits depend on the value of the fund F(t).
- Minimal annual return  $\kappa_g$  and maximal annual return  $\kappa_m$  on F(t) with participation rate  $\eta$ .
- The survival benefit is given by  $F_T^e \times \mathbb{I}\{t \geq T\}$  and the death benefit by  $F_t^e \times \mathbb{I}\{t < T\}$  where :

$$F_t^e = F_0 \prod_{u=1}^{\lfloor t \rfloor} \min \left\{ e^{\kappa_m} \; ; \; \max \left\{ 1 + \eta \left( \frac{S_u}{S_{u-1}} - 1 \right) \; ; \; e^{\kappa_g} \right\} \right\}$$

The faire value is given by discounting the expected benefits under a risk-neutral measure  $\mathbb{P}^*$  with mortality modeled by a Poisson process (Makeham force of mortality  $\mu_x$ ).

## FAIR VALUES OF LIFE INSURANCE CONTRACTS

- 50-year-old female policyholder,  $F(0) = 10\,000 \in$  ,  $\kappa_m = 20\%$  et  $\eta = 80\%$ .
- Fair value  $FV_0$  for different maturities T with  $\kappa_g=1\%$  :

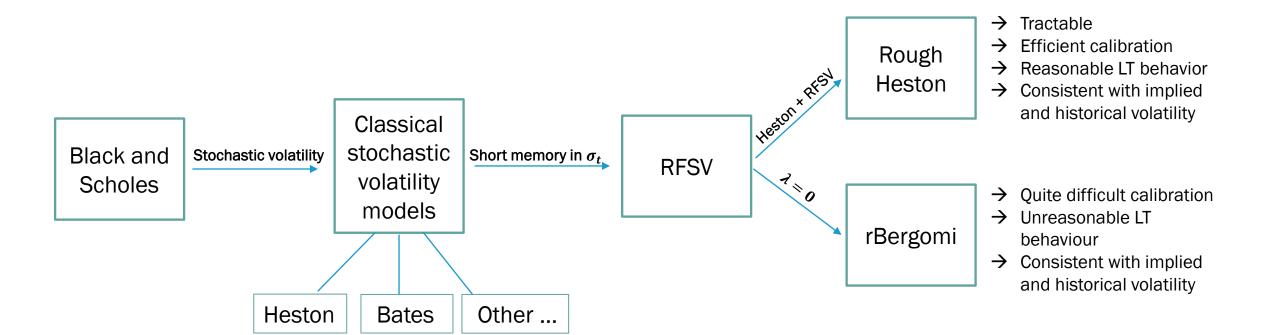
|             |             |             |             |   |             | _                         |
|-------------|-------------|-------------|-------------|---|-------------|---------------------------|
| $k_g = 1\%$ | Heston      | Bates       | rBergomi    | ſ | ough Heston |                           |
| T=5         | 14 036.57 € | 13 457.49 € | 12 291.17 € |   | 13 529.65 € | Most market-consistent    |
| T = 10      | 19 372.58 € | 17696.27 €  | 14 909.54 € |   | 14 063.09 € | and accurate fair values! |
| T = 20      | 35 631.76 € | 29 613.34 € | 20 639.81 € |   | 17823.38 €  |                           |
|             |             |             |             |   |             | _                         |

- → Lower fair values of rough-type models compared with the Heston and Bates models.
- → Higher fair values of the **rBergomi** model compared with the **rough Heston** model for **large maturities** *T* (non-stationarity).

### WHY USING THE ROUGH HESTON MODEL?

- The <u>rough Heston</u> model allows :
- To better reproduce the observed historical volatility.
- A better modeling of the **implied volatility surface** (ATM volatility skews).
- An easy, efficient and stable calibration method with only 3 parameters.
- Reasonable long-term properties due to its **stationary volatility** process.
  - → The <u>rough Heston</u> tends to outperform existing models in terms of long-term pricing of insurance contrats

# TO SUMMARIZE ...



Thank you for your attention!